

# Lecture 26

## Hypothesis Testing

---

Manju M. Johnny

STAT 330 - Iowa State University

1 / 11

## Hypothesis Testing

### Definition:

A statistical *hypothesis* is a statement about a parameter  $\theta$

There are 2 competing hypotheses in a testing problem:

- *Null Hypothesis ( $H_0$ )*: the default/pre-data view about the parameter. (What we already believe - never prove this)
- *Alternative Hypothesis ( $H_A$ )*: usually what you want your data/study to show. (what you're trying to prove)

**Note:**  $H_0$  and  $H_A$  have to be disjoint. There can not be any outcomes in common between the null and alternative hypotheses.

2 / 11

## Motivating Example

Example 1: I have a coin and I'm interested in the probability of flipping a "head". I flip a coin 100 times and record the number of heads obtained.

$$X = \# \text{ of heads}$$

$$X \sim \text{Bin}(n = 100, p)$$

where  $p = P(\text{"heads"})$  is unknown

By default, we assume coin is fair  $p = 0.5$  (null hypothesis).

Alternative hypothesis should contradict the null hypothesis.

### Hypotheses:

- $H_0 : p = 0.5$  (coin is fair)
- $H_A : p \neq 0.5$  (coin is unfair)

3 / 11

## Motivating Example Continued

Data: Out of 100 flips, I get 71 heads.  $\hat{p} = 0.71$

### Idea of Hypothesis Testing:

- Assume  $H_0$  (our default belief) is true until our *data* tells us otherwise.
- Ask ourselves "what is the probability of getting 71 heads if the null hypothesis is true (coin is fair)?"  
→ probability = 0.000032 (called the "*p - value*")
- There is a 0.000032 probability that we observed our data if the null hypothesis that the coin is fair is true.  
→ Now we have evidence against the null hypothesis (that coin is fair), and in favor of the alternative hypothesis (that coin is unfair).

4 / 11

## General Hypothesis Testing Procedure

---

### Hypothesis Tests

We will look at 4 different hypothesis testing scenarios.

Their null hypotheses are given below:

- $H_0 : \mu = \#$
  - $H_0 : p = \#$
  - $H_0 : \mu_1 - \mu_2 = \#$
  - $H_0 : p_1 - p_2 = \#$
- usually 0
- } test whether the population parameter is equal to some value (#).
- } test whether the difference in parameters for 2 groups is equal to some value (#). This value (#) is almost always 0.

The above all follow the same general hypothesis testing procedure.

# Testing Procedure

## General Hypothesis Testing Procedure

1. Determine the Null and Alternative Hypotheses:
- always "=" →  $H_0: \theta = \#$   
sign depends →  $H_A: \theta < \#$   
on problem  $\quad \quad \quad > \#$   
 $\quad \quad \quad \neq \#$

# should be same in  $H_0$  and  $H_A$

Note:  $\theta$  is just stand in symbol for parameter  $\theta$  can be  $\begin{cases} \mu \\ p \\ \mu_1 - \mu_2 \\ p_1 - p_2 \end{cases}$

2. Gather data and calculate a **test statistic** under the assumption that  $H_0$  is true. Test statistic has general form:

$$Z = \frac{\hat{\theta} - \#}{SE(\hat{\theta})}$$

3. Calculate the **p-value**. Use p-value to determine whether you have enough evidence to reject the null hypothesis.
- small p-value →  $H_0$  unlikely → Reject  $H_0$
  - large p-value → No evidence against  $H_0$  → Do not reject  $H_0$

Fail to reject  $H_0$  6/11

## Calculating p-values

## Calculating $p$ -value

### Definition: $p$ -value

The  $p$ -value is the probability of observing your test statistic or more extreme if the null hypothesis ( $H_0$ ) is true.

"more extreme" can be bigger, smaller or both depending on the the sign in the alternative hypothesis ( $H_A$ )

- Small  $p$  - value indicates a small probability of seeing your data if  $H_0$  is true. The data is evidence against  $H_0$  (Reject  $H_0$ )
- Large  $p$  - value indicates a large probability of seeing your data if  $H_0$  is true. No evidence against  $H_0$  (Do Not Reject  $H_0$ )
- $P$  - value is often *wrongly* interpreted as the probability of the null hypothesis. (Don't make this mistake)

7 / 11

## Calculating the $p$ - value

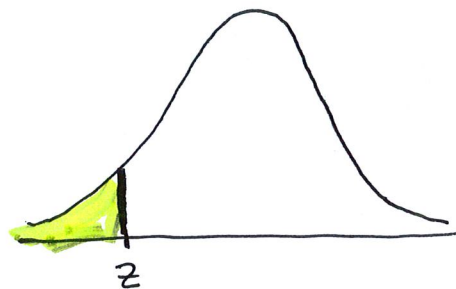
- By central limit theorem, the estimator follows a normal distribution. Standardizing the estimator gives us the test statistic  $Z$ , which follows  $N(0, 1)$  distribution
- Obtain  $p$  - value from the  $z$ -table as left-hand area, right-hand area or both (depending on sign in  $H_A$ )

### Left-sided Hypothesis Test

$$H_0 : \theta = \#$$

$$H_A : \theta < \#$$

$$Z = \frac{\hat{\theta} - \#}{SE(\hat{\theta})}$$



$$p\text{-value} = P(Z \leq z)$$

$z$   
test statistic  
value  
8 / 11

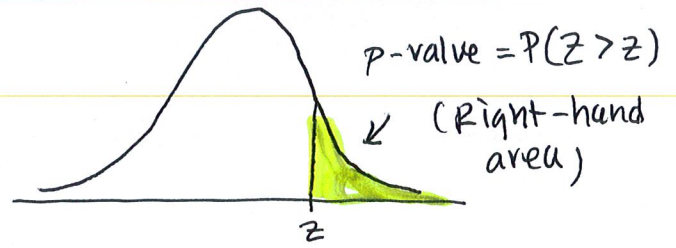
## Calculating p-value Cont.

### Right-sided Hypothesis Test

$$H_0: \theta = \#$$

$$H_A: \theta > \#$$

$$Z = \frac{\hat{\theta} - \#}{SE(\hat{\theta})}$$

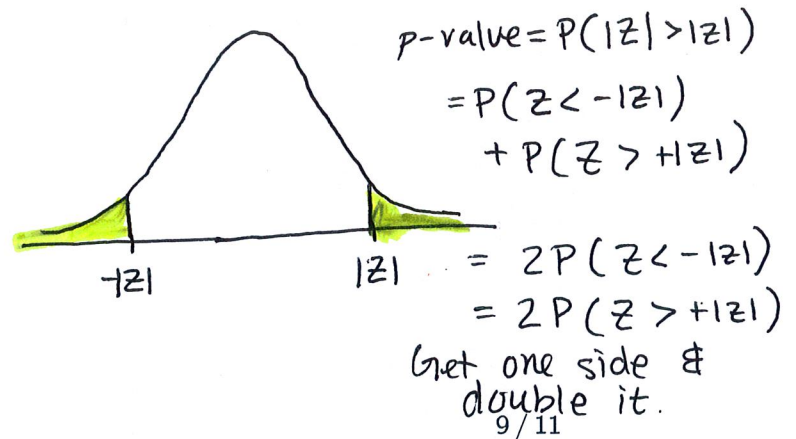


### 2-sided Hypothesis Test

$$H_0: \theta = \#$$

$$H_A: \theta \neq \#$$

$$Z = \frac{\hat{\theta} - \#}{SE(\hat{\theta})}$$



Easiest way: make your  $z$  a negative value, and find the left-hand area, then double it.

## Types of Errors

In the testing framework, it is possible to make errors that are inherent to the testing procedure (not calculation mistakes).

### Types of errors

- Type I Error (wrongly reject  $H_0$ )  
 $\rightarrow P(\text{Type I error}) = \alpha$
- Type II Error (wrongly fail to reject  $H_0$ )  
 $\rightarrow P(\text{Type II error}) = \beta$

		Decision	
		Reject $H_0$	Don't Reject $H_0$
Truth	$H_0$ True	Type I error	No error
	$H_0$ False	No Error	Type II error

### Note:

- $\alpha$  (significance level) can be viewed as a cut-off for how small the  $p$ -value needs to be to reject  $H_0$ . Reject  $H_0$  if  $p\text{-value} < \alpha$ . ( $\alpha$  set before conducting the test).
- In this class, we use a strength of evidence argument without a "cut-off" for  $p$ -value.

Not on Final

## Hypothesis Testing Examples

### Tax Fraud Example

#### Example: Tax Fraud

Default  
belief →

Historically, IRS taxpayer compliance audits have revealed that about 5% of individuals do things on their tax returns that invite criminal prosecution.

A sample of  $n = 1000$  tax returns produces  $\hat{p} = 0.061$  as an estimate of the fraction of fraudulent returns.

Does this provide a clear signal of change in the tax payer behavior?

↳ no specific direction.  
~~we~~ can be greater or less  
use " $\neq$ " in  $H_A$

1. State the Hypotheses

$$H_0 : p = 0.05$$

$$H_A : p \neq 0.05$$

↪ same  
value  
(never use  
 $\hat{p}$ )

## Tax Fraud Example

2. The *test statistic* will be obtained from

$$Z = \frac{\hat{p} - \#}{\sqrt{\frac{\#(1-\#)}{n}}} = \frac{\hat{p} - 0.05}{\sqrt{\frac{0.05(0.95)}{n}}}$$

Under the null hypothesis,  $Z$  follows a  $N(0,1)$  distribution.

data Plugging in our data values, we get the test statistic

$$\hat{p} = 0.061$$
$$n = 1000$$

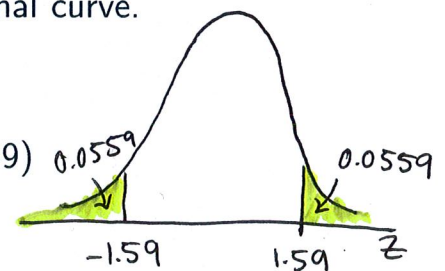
$$z = \frac{0.061 - 0.05}{\sqrt{\frac{0.05(0.95)}{1000}}} = 1.59$$

2/9

## Tax Fraud Cont.

3. Since we have a " $\neq$ " in the  $H_A$ , the *p-value* is obtained from both the *left-hand* and *right-hand area* of the normal curve.

$$\begin{aligned} p\text{-value} &= P(|Z| \geq 1.59) \\ &= P(Z < -1.59) + P(Z > 1.59) \\ &= 2 \cdot P(Z < -1.59) \\ &= 2 * 0.0559 \\ &= 0.1118 \quad \approx 11.18\% \end{aligned}$$



$$\Rightarrow H_0: p = 0.05$$
$$\chi H_A: p \neq 0.05$$

This is not a very small *p*-value. We therefore only have very weak evidence against  $H_0$ . Thus, we *do not* reject the null hypothesis in favor of the alternative hypothesis.

There is not much evidence of change in tax payer behavior.

3/9



## Disk Drive Example

Example: Disk Drive

$n_1 = 30$  and  $n_2 = 40$  disk drives of 2 different designs were tested under conditions of "accelerated" stress and times to failure recorded:

	Group 1	Group 2
	Standard Design	New Design
$n_1 = 30$		$n_2 = 40$
$\bar{x}_1 = 1205$ hr		$\bar{x}_2 = 1400$ hr
$s_1 = 1000$ hr		$s_2 = 900$ hr

Does the new design have a larger mean time to failure under "accelerated" stress? In other word, is the new design better?

1. State the Hypotheses

$$H_0 : \mu_1 = \mu_2 \quad \Rightarrow \quad \mu_1 - \mu_2 = 0$$

$$H_A : \mu_1 < \mu_2 \quad \Rightarrow \quad \mu_1 - \mu_2 < 0$$

4 / 9

## Disk Drive Cont.

2. The *test statistic* will be obtained from

$$Z = \frac{(\bar{X}_1 - \bar{X}_2) - 0}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$$

Under the null hypothesis,  $Z$  follows a  $N(0,1)$  distribution.

Plugging in our data values, we get the test statistic

$$z = \frac{(1205 - 1400) - 0}{\sqrt{\frac{1000^2}{30} + \frac{900^2}{40}}} = -0.84$$

↑  
observed  
test  
statistic

5 / 9

## Disk Drive Cont.

3. Since we have a " $<$ " in the  $H_A$ , the  $p$ -value is obtained from the **left-hand area** of the normal curve.

$$\begin{aligned} p\text{-value} &= P(Z < -0.84) \\ &= 0.2005 \quad \approx 20\% \end{aligned}$$

$$\Rightarrow H_0: \mu_1 = \mu_2$$

$$\times H_A: \mu_1 < \mu_2$$

This is not a small  $p$ -value. We therefore only have very weak evidence against  $H_0$ . Thus, we do not reject the null hypothesis in favor of the alternative hypothesis.

There is not significant evidence that the new design is better.

6/9

## Queuing System Example

### Example: Queuing System

Suppose we have 2 queuing systems A and B. We'd like to know whether system A has a higher probability of having an available server in the long run than system B. The simulation data for the 2

servers is shown below:

Group 1	Group 2
System A	System B
$n_1 = 500$ runs	$n_2 = 1000$ runs
$\hat{p}_1 = \frac{303}{500} = 0.606$	$\hat{p}_2 = \frac{551}{1000} = 0.551$

$\hat{p}$  = sample proportion of available server in runs

where  $\hat{p}$  is the proportion runs with available servers at  $t = 2000$ .

1. State the Hypotheses

$$H_0: p_1 = p_2 \quad \Rightarrow \quad p_1 - p_2 = 0$$

$$H_A: p_1 > p_2 \quad \Rightarrow \quad p_1 - p_2 > 0$$

7/9

## Queuing System Cont.

2. The *test statistic* will be obtained from

$$Z = \frac{(\hat{p}_1 - \hat{p}_2) - 0}{\sqrt{\hat{p}_{pool}(1 - \hat{p}_{pool})} \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

Under the null hypothesis,  $Z$  follows a  $N(0,1)$  distribution.

Next, calculate  $\hat{p}_{pool}$  to plug into the denominator of the test statistic.

$$\hat{p}_{pool} = \frac{n_1 \hat{p}_1 + n_2 \hat{p}_2}{n_1 + n_2} = \frac{303 + 551}{500 + 1000} = 0.569$$

Plugging in our data values, we get the test statistic

$$z = \frac{(0.606 - 0.551) - 0}{\sqrt{0.569(1 - 0.569)} \sqrt{\frac{1}{500} + \frac{1}{1000}}} = 2.03$$

observed  
test statistic  
value

8/9

## Queuing System Cont.

3. Since we have a " $>$ " in the  $H_A$ , the *p-value* is obtained from the right-hand area of the normal curve.

$\times H_0: P_1 = P_2$   
 $\Rightarrow H_A: P_1 > P_2$

$$\begin{aligned} p\text{-value} &= P(Z > 2.03) \\ &= 1 - 0.9788 \\ &= 0.0212 \approx 2\% \end{aligned}$$

This is a small  $p$ -value. We therefore have strong evidence against  $H_0$ . Thus, we reject the null hypothesis in favor of the alternative hypothesis.

There is strong evidence that system A has a higher probability of having an available server than system B.

9/9

# Hypothesis Testing Summary

Null Hypothesis	Test-Statistic	Reference Dist.
$H_0 : \mu = \#$	$Z = \frac{\bar{X} - \#}{s/\sqrt{n}}$	$Z \sim N(0, 1)$
$H_0 : p = \#$	$Z = \frac{\hat{p} - \#}{\sqrt{\frac{\#(1-\#)}{n}}}$	$Z \sim N(0, 1)$
$H_0 : \mu_1 - \mu_2 = \#$	$Z = \frac{(\bar{X}_1 - \bar{X}_2) - \#}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$	$Z \sim N(0, 1)$
$H_0 : p_1 - p_2 = \#$	$Z = \frac{(\hat{p}_1 - \hat{p}_2) - \#}{\sqrt{\hat{p}_{pool}(1 - \hat{p}_{pool})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$ where $\hat{p}_{pool} = \frac{n_1\hat{p}_1 + n_2\hat{p}_2}{n_1 + n_2}$	$Z \sim N(0, 1)$

"#" is usually 0